Traditional and modern psychometrics using R: Presented as part of a symposium on Using R in personality research A symposium at the First World Conference on Personality

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Exploratory Factor Analysis

Classical and IRT approaches to test construction 000000000000

William Revelle : Traditional and modern psychometrics

Traditional and modern psychometrics using R

William Revelle Northwestern University

To understand how personality constructs relate to other psychological constructs as well as real world criteria, it is first necessary to develop and assess the reliability and validity of personality scales. The psych package in R has been developed for this purpose. Basic item analysis, factor and cluster analysis, reliability analysis, and item response measures can be done in the psych (Revelle, 2012) package. Two other packages, sem (Fox, 2012) and lavaan (Rosseel, 2012) have been developed to allow confirmatory factor analysis and structural equation modeling. The use of all three of these packages in measuring and evaluating the structure of personality constructs will be demonstrated.



What is the structure of mood?



- Positive and Negative Affect (Watson & Tellegen, 1985)
- Valence and Arousal (Barrett & Russell, 1998)
- Energetic and Tense Arousal (Thayer, 1978, 2000)
- Various psychometric solutions
 - Two dimensional simple structure models
 - Two dimensional circumplex models
- Various problems
 - Unipolar vs. bipolar items (Russell & Carroll, 1999)
 - Item skew (Rafaeli & Revelle, 2006)



Analysis of the Motivational State Questionnaire

The Motivational State Questionnaire (MSQ) was developed to study mood and arousal (Revelle & Anderson, 1997). It included adjectives from a variety of sources. Included in *psych* as the msq data set are 75 mood and arousal items given to 3896 subjects over 10 years.

```
> f2 <- fa(msq[1:70],2)
```

```
> text(f2$loadings,
      rownames(f2$loadings),cex=.5)
```

- Factor analyze the first 70 msq items. Extract two factors.
- Plot the resulting solution, setting the size of the x and y axes. Use a small plot character.
- Add labels for each data point. Use a small character size.

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Conclusion

Structure of Mood

Structure of MSQ emotions using Pearson R

Circumplex of emotions





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Conclusion

Structure of Mood

Structure emotion using polychoric correlations

Because the MSQ items were 1-4 we should not treat them as continuous, but rather as categorical. We can find polychoric correlations to compensate for skew.

```
> msqR <- polychoric(msq[1:70])</pre>
```

```
> f2p <- fa(msqR$rho,2)
```

- Find the polychroric correlations of the first 70 msq items
- Factor analyze the resulting correlations. Extract two factors.
- Out the resulting solution, setting the size of the x and y axes. Use a small plot character.
- Add labels for each data point. Use a small character



Structure of MSQ emotions using polychoric R

Circumplex of emotions using polychoric r





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Compare with the structure of MSQ emotions using Pearson R

Circumplex of emotions





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Compare the 2 solutions in terms of factor congruence

Factor congruence is just the cosine of the angle between the factors:

$$r_{c} = \frac{\sum_{1}^{n} F_{xi} F_{yi}}{\sqrt{\sum_{1}^{n} F_{xi}^{2} \sum_{1}^{n} F_{yi}^{2}}}.$$

or

 $diag(\mathbf{F_xF_x}')^{-1/2}$

It may be found using the factor.congruence function. We should not just correlate the loadings.

 MR1
 MR2
 MR1
 MR2

 MR1
 1.00
 -0.04
 MR1
 1.00
 -0.40

 MR2
 -0.06
 1.00
 MR2
 -0.39
 0.99

 The factors are essentially identical.
 MR2
 -0.39
 0.99



Factor Extension and Set Correlation as ways of relating multiple domains

Factor Extension and Set Correlation

- Originally developed by Dwyer (1937) for the case of having completed a factor analysis and then a new variable is introduced.
 - At the time, factoring was hard and time consuming
- May now be used to extend the factors from one domain into another domain (Horn, 1973).
 - Differs from SEM in that the factors are estimated in the first domain and are not changed with the addition of the second domain
- Another technique for relating two domains is "Set Correlation" as discussed by Cohen, Cohen, West & Aiken (2003)



Factor Extension and Set Correlation as ways of relating multiple domains

Factor Extension and the structure of affect

- Consider the joint analysis of Energetic and Tense Arousal with Positive and Negative Affect
 - EA = "active" "alert" "aroused" ("sleepy" "tired" "drowsy")
 - TA = "anxious" "jittery" "nervous" -("calm" "relaxed" "at-ease")
 - PA = "happy" "pleased"
 - NA = "unhappy" "sad"
- ② What is the location of PA and NA in the EA/TA space
- What is the structure of the joint space?
- Use the data in the Motivational State Questionnaire (msq) data set.
 - 75 mood and arousal items given over 10 years to various participants (N=3896)



Factor Extension and Set Correlation as ways of relating multiple domains

Basic commands for display and and analysis

```
eata <- c("active","alert","aroused",
"sleepy","tired","drowsy",
"anxious","jittery","nervous",
"calm","relaxed","at-ease",
"happy","pleased","unhappy","sad")</pre>
```

```
R <- lowerCor(msq[eata])</pre>
```

cor.plot(R,main="Arousal and Affect terms")

f.all <- fa(R,2)
fe.all <- fa.extend(R,2,1:12,13:16)</pre>



- Ind the correlations
- show the correlations graphically
- 4 factor entire set
- factor EA/TA space extend to PA/NA



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Factor Extension and Set Correlation as ways of relating multiple domains

A cor.plot of the data

Arousal and Affect terms





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Factor Extension and Set Correlation as ways of relating multiple domains

fa(r = R, nfactors = 2)fa.extend(r = R, nfactors = 2, ov = 1:12, ev = 13:16)Factor Analysis using method = minres Factor Analysis using method = minres Call: fa(r = R, nfactors = 2)Call: fa.extend(r = R, nfactors = 2, ov = 1:12, ev = 1) Standardized loadings (pattern matrix) Standardized loadings (pattern matrix) MR1 MR2 h2 u2 MR1 MR2 h2 u2 active -0.52 0.25 0.39 0.61 active -0.57 0.02 0.32 0.68 alert -0.64 0.22 0.52 0.48 alert -0.68 0.07 0.47 0.53 aroused -0.46 0.16 0.27 0.73 aroused -0.49 -0.07 0.24 0.76 sleepy 0.89 0.06 0.78 0.22 sleepy 0.88 0.01 0.78 0.22 tired 0.86 0.01 0.73 0.27 tired 0.85 -0.01 0.73 0.27 drowsv 0.88 0.07 0.75 0.25 drowsv 0.87 0.01 0.76 0.24 anxious -0.21 -0.34 0.13 0.87 anxious -0.14 -0.50 0.26 0.74 jittery -0.31 -0.34 0.17 0.83 jittery -0.23 -0.53 0.33 0.67 nervous -0.15 -0.40 0.16 0.84 nervous -0.07 -0.55 0.30 0.70 calm 0.18 0.67 0.43 0.57 calm0.04 0.68 0.46 0.54 relaxed 0.07 0.71 0.48 0.52 relaxed -0.08 0.69 0.49 0.51 at-ease 0.00 0.74 0.55 0.45 at-ease -0.15 0.69 0.51 0.49 happy -0.30 0.59 0.51 0.49 happy -0.49 0.32 0.36 0.64 pleased -0.28 0.53 0.42 0.58 pleased -0.45 0.27 0.29 0.71 unhappy 0.14 -0.45 0.25 0.75 unhappy 0.22 -0.36 0.19 0.81 0.11 -0.39 0.19 0.81 0.17 -0.33 0.15 0.85 sad sad MR1 MR2 MR1 MR2 SS loadings SS loadings 3.65 3.07 3.95 2.69 Proportion Var Proportion Var 0.23 0.19 0.25 0.17 Cumulative Var Cumulative Var 0.23 0.42 0.25 0.42 Proportion Explained 0.54 0.46 Proportion Explained 0.59 0.41 Cumulative Proportion 0.54 1.00 Cumulative Proportion 0.59 1.00 With factor correlations of With factor correlations of MR2 MR1 MR2 MR1 MR1 1.00 -0.21 MR1 1.00 -0.06 MR2 -0.21 1.00 MR2 -0.06 1.00



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Conclusion

Factor Extension and Set Correlation as ways of relating multiple domains

A fa.plot of the two solutions



Factor Extension and Set Correlation as ways of relating multiple domains

Factor extension of Energetic and Tense Arousal to Affect

EA and TA factors extended to PA and NA





Factor Extension and Set Correlation as ways of relating multiple domains

Set correlation is a generalized R^2 between two sets of variables

 $R^2 = 1 - \prod (1 - \lambda_i^2)$ where λ_i^2 is the is ith squared canonical correlation. Unfortunately, the R^2 is sensitive to one of the canonical correlations being very high. An alternative, T^2 , is the proportion of additive variance and is the average of the squared canonicals (Cohen et al., 2003).

> set.cor(y=13:16,x=1:12,data=R)

Call: set.cor(y = 13:16, x = 1:12, data = R)

Multiple Regression from matrix input

Beta weights

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	happy	pleased	unhappy	sad
active	0.28	0.25	-0.07	-0.02
alert	0.17	0.15	0.05	0.01
aroused	0.16	0.20	-0.05	-0.04
sleepy	0.04	0.05	0.03	0.08
tired	-0.03	-0.05	0.17	0.14
drowsy	0.01	0.03	0.00	-0.04
anxious	0.01	0.01	0.10	0.17
jittery	0.02	0.00	-0.04	-0.03
nervous	-0.01	0.01	0.19	0.20
calm	0.08	0.08	0.00	0.04
relaxed	0.13	0.10	-0.10	-0.06
at-ease	0.20	0.17	-0.12	-0.10

> set.cor(y=13:16,x=1:12,data=R)

Multiple	e R		
happy	pleased	unhappy	sad
0.69	0.64	0.43	0.41

Multiple R2 happy pleased unhappy sad 0.47 0.41 0.18 0.17

Various estimates of between set correlations Squared Canonical Correlations [1] 0.5187 0.1551 0.0095 0.0041 Chisq of canonical correlations NULL

Average squared canonical correlation = 0.17Cohen's Set Correlation R2 = 0.6

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Classical Reliability Estimates

- Guttman (1945) considered 6 different estimates of reliability. Of these, one λ_3 is also known as α (Cronbach, 1951).
- ⁽²⁾ McDonald (1999) introduced two additional reliability coefficients β which we (Zinbarg, Revelle, Yovel & Li, 2005; Revelle & Zinbarg, 2009) refer to as $\omega_{hierarchical}$ and ω_{total} .
 - $\omega_{hierarchical}$ or ω_h is an estimate of the general factor saturation of a test.
 - ω_{total} or ω_t is an estimate of the total reliable variance in a test.
- ③ All of these estimates of reliability are available in the *psych*.
 - $\alpha\,$ alpha, guttman, omega, score.items
 - λ_{1-6} guttman
 - $\omega_h, \, \omega_t$ omega



α can be misleading if applied to multifactorial items

Score the two dimensions of the Energetic and Tense Arousal items as one scale

$\lambda = \frac{1}{2}$			Item statistics			
> alpha(ms)	q[eata[1:12	2]])			n rr.corr.drop mean sd	
Reliabilit	v analysis				active 3890 0.73 0.70 0.627 1.03 0.93	
Call: alph	a(x = msolution)	0ata [1·12	ווו		alert 3885 0.78 0.77 0.714 1.15 0.91	
ourr. arph					aroused 3890 0.66 0.62 0.543 0.71 0.85	
raw alph	a std alph	a C6(smc)	average	r meen sd	sleepy- 3880 0.69 0.71 0.620 1.25 1.05	
	a sta.arpn				tired- 3886 0.70 0.70 0.629 1.39 1.04	
0.76 0.74 0.84 0.19 1.		1.1 0.52	drowsy- 3884 0.67 0.68 0.600 1.16 1.03			
Bolishili	ty if an i	tom is dr	opped.		anxious 2047 0.33 0.24 0.134 0.67 0.86	
nerrabiti	u alpha st	d alpha C	oppeu. 6(smc) au	erage r	jittery 3890 0.37 0.29 0.189 0.59 0.80	
activo	∞ arpina SC			0 17	nervous 3879 0.25 0.16 0.066 0.35 0.65	
accive	0.71	0.03	0.02	0.17	calm 3814 0.23 0.15 0.084 1.55 0.92	
arent	0.70	0.08	0.01	0.10	relaxed 3889 0.32 0.25 0.190 1.68 0.88	
aloopu-	0.73	0.70	0.82	0.18	at-ease 3879 0.41 0.36 0.283 1.59 0.92	
sieepy-	0.71	0.70	0.01	0.17		
drougu-	0.71	0.70	0.81	0.17	Non missing response frequency for each item	
arowsy-	0.72	0.70	0.81	0.10	0 1 2 3 miss	
anxious	0.77	0.75	0.65	0.21	active 0.35 0.35 0.23 0.07 0.00	
jittery	0.76	0.74	0.84	0.21		
nervous	0.77	0.76	0.85	0.22	anxious 0.55 0.28 0.13 0.04 0.47	
caim	0.78	0.76	0.85	0.22		
relaxed	0.77	0.75	0.84	0.21	calm 0.13 0.35 0.35 0.17 0.02	
at-ease	0.76	0.74	0.84	0.20	relaxed 0.10 0.31 0.41 0.18 0.00	
					at-ease 0.13 0.33 0.37 0.17 0.00	
					Warning message:	
					<pre>In alpha(msq[eata[1:12]]) :</pre>	
					Some items were negatively correlated with $_{25}$; / 73
					total scale and were automatically reversed \sim	/ 13

Compare α to ω_h for this multifactorial set of items

> omega(msq[eata[1:12]],2)

Omega

Ο

Call:	<pre>omega(m = msq[</pre>	eata[1:12
Alpha:		0.75
G.6:		0.85
Omega	Hierarchical:	0.09
Omega	H asymptotic:	0.11
Omega	Total	0.83

Schmid Leiman Factor loadings greater than 0.2					0.2		
	g	F1*	F2*	h2	u2	p2	
active-		0.55		0.32	0.68	0.06	
alert-		0.66		0.47	0.53	0.07	
aroused-		0.48		0.24	0.76	0.04	
sleepy	0.21	0.86		0.78	0.22	0.06	
tired	0.20	0.83		0.73	0.27	0.06	
drowsy	0.20	0.85		0.76	0.24	0.05	
anxious			-0.48	0.26	0.74	0.03	
jittery		-0.23	-0.52	0.33	0.67	0.02	
nervous			-0.53	0.30	0.70	0.04	
calm-			-0.66	0.46	0.54	0.05	
relaxed-			-0.67	0.49	0.51	0.07	
at-ease-	0.20		-0.67	0.51	0.49	0.08	
With organization of							

with eigenvalues of: g F1* F2*

0.31 3.22 2.13

^{2]], nfactors} $(1)^{2}\omega_h$ is a higher order factor model and requires 3 lower level factors for identification.

- It can be found with two factors under various assumptions.
- By default, omega assumes equal loadings of the lower level factors on the higher order factor, but this may be changed.
- A warning is given for this condition.



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Representing a higher order structure

- Omega_h may be found by Exploratory Factor Analysis by factoring the data, applying an oblique transformation (e.g., oblimin) and then factoring the correlation matrix of these resulting factors. Factor loadings on the general factor are then found using the Schmid & Leiman (1957) transformation.
- Alternatively, omegah may be directly estimated using Confirmatory Factor Analysis using the sem (Fox, Nie & Byrnes, 2012) or lavaan (Rosseel, 2012) packages.
- Image of the solution of the solution.
- The graphical representation of the Schmid-Leiman transformation is automatically drawn by omega.



This shows that there is no general factor of these two dimensions

Omega with Schmid Leiman Transformatio

Hierarchical (multilevel) Structure







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Consider another example: 16 ability items

- **1** 16 ability items reflecting 4 subdomains for N=1525.
- 2 Example is taken from iqitems in *psych*.
- Collected using SAPA (Synthetic Aperture Personality Assessment) as part of the ICAR (International Cognitive Ability Resource) project.
- Convert multiple choice to Correct/Incorrect
- Score for traditional α using alpha as well as ω_h .



Finding α and ω_h for 16 ability items

> data(iqitems)

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- > iq.keys <- c(4,4,4, 6, 6,3,4,4, 5,2,2,4, 3,2,6,7)
- > score.multiple.choice(iq.keys,iqitems)
- > iq.scrub <- scrub(iqitems,isvalue=0)</pre>
- > iq.tf <- score.multiple.choice(
 iq.keys,iq.scrub, score=FALSE)</pre>
- > alpha(iq.tf)
- > omega(iq.tf,nfactors=4)

- Get the data
- Assign a scoring key
- Score the items to get summary statistics
- Convert non-responses to missing (NA)
- Convert the multiple choice items to correct/incorrect
- **(5)** Find conventional α
- Find ω_h



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Classical test theory – going beyond α

Comparing α and ω_h for hierarchically organized data

> omega(iq.tf,nfactors=4)

	Omega
<pre>> alpha(ig.tf)</pre>	Call: omega(m = iq.tf, nfactors = 4)
	Alpha: 0.83
Reliability analysis	G.6: 0.84
Call: $alpha(x = ig.tf)$	Omega Hierarchical: 0.65
	Omega H asymptotic: 0.76
raw_alpha std.alpha G6(smc) average_r mean sd	Omega Total 0.86
0.83 0.83 0.84 0.23 0.49 0.25	5
	Schmid Leiman Factor loadings greater than 0.2
Reliability if an item is dropped:	g F1* F2* F3* F4* h2 u2 p2
raw alpha std.alpha G6(smc) average r	reason.4 0.50 0.27 0.34 0.66 0.73
reason.4 0.82 0.82 0.82 0.23	reason.16 0.42 0.21 0.23 0.77 0.76
	reason.17 0.55 0.47 0.52 0.48 0.57
rotate.8 0.82 0.82 0.83 0.24	reason.19 0.44 0.21 0.25 0.75 0.77
	letter.7 0.52 0.35 0.39 0.61 0.69
Item statistics	letter.33 0.46 0.30 0.31 0.69 0.70
n rr.corr.drop mean sd	letter.34 0.54 0.38 0.43 0.57 0.67
reason.4 1442 0.58 0.54 0.50 0.68 0.47	letter.58 0.47 0.20 0.28 0.72 0.78
	matrix.45 0.40 0.66 0.59 0.41 0.27
rotate.8 1460 0.51 0.47 0.41 0.19 0.39	matrix.46 0.40 0.26 0.24 0.76 0.65
	matrix.47 0.42 0.23 0.77 0.79
Non missing response frequency for each item	matrix.55 0.28 0.12 0.88 0.65
0 1 miss	rotate.3 0.36 0.61 0.50 0.50 0.26
reason.4 0.32 0.68 0.05	rotate.4 0.41 0.61 0.54 0.46 0.31
	rotate.6 0.40 0.49 0.41 0.59 0.39
rotate.8 0.81 0.19 0.04	rotate.8 0.32 0.53 0.40 0.60 0.26
	With eigenvalues of: 11/73
	σ F1* F2* F3* F4*

Bifactor solution to the 16 ICAR ability items shows g and first order factors

Bifactor structure of 16 ICAR cognitive ability items





IRT measures of reliability

2 parameter IRT is equivalent to EFA solution

- Item Response Theory approaches consider item difficulty and item discrimination.
 - 1 parameter IRT considers just item location and applies the Rasch model. Can be found using the *ltm* package.
 - 2 parameters of IRT are location and discrimination. These are reparameterizations of factor loadings and item difficulty: That is, 2 parameter IRT models are just factor models applied to the *tetrachoric* or *polychoric* correlations.
 - That is, find the factor analysis loadings (λ_i) and the item endorsement frequencies expressed as normal deviates (τ_i and then convert to IRT parameters
 - discrimination $\alpha = \frac{\lambda_i}{\sqrt{1-\lambda_i^2}}$

• location (difficulty)
$$\overset{\mathbf{v}}{\delta} = \frac{\tau_i}{\sqrt{1-\lambda_i^2}}$$

IRT statistics can be done using irt.fa.



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IRT measures of reliability

Item information for a 1 factor solution

Item information from factor analysis



IRT measures of reliability

Test information for a 1 factor solution

Test information -- item parameters from factor analysis



IRT measures of reliability

Item information for each lower level factor of 16 ICAR items



Latent Trait (normal scale)

Item information from factor analysis

Item information from factor analysis



Latent Trait (normal scale)

Item information from factor analysis

Item information from factor analysis



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Using R for personality research: Classical and Modern psychometrics

Combining the power of base R with additional packages allows personality researchers to

- Do basic scale construction
- 2 Perform classical (α) and more advanced (ω_h, ω_t) analyses of reliability.
- Operation Perform Exploratory and Confirmatory Factor Analysis
- O "modern" psychometrics using Item Response Theory

